International Journal of Mathematics and Computer Science Volume **20**, Issue no. 1, (2025), 387–391 DOI: https://doi.org/10.69793/ijmcs/01.2025/khan

### Generalized Order Statistics from a Novel Family of Distribution

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(Received June 6, 2024, Accepted July 7, 2024, Published November 25, 2024)

#### Abstract

In this article, we establish the moment's properties of a novel family of distribution are via generalized order statistics (GOS). We take into consideration the cases for order statistics and record values.

## 1 Introduction

In 1995, the probability distributions have been investigated by order statistics and record values separately. Kamps [1] proposed a model that unifies several models (order, sequential order, record, and Pfeifer's record) into one frame called GOS. The mechanism of this frame is arranged in ascending order of random variables.

Let  $X_1, X_2, \dots, X_n$  be a sequence of independent and identically distributed (IID) random variables (*RV*) based on distribution function (*DF*) F(x) and probability density function (*PDF*) f(x). Then the joint *PDF* of *GOS* is

$$f_{r,n,\tilde{m},k}(x_1,\cdots,x_n) = k\left(\prod_{j=1}^{n-1}\gamma_j\right) \left[\bar{F}(x_n)\right]^{k-1} f(x_n)\left(\prod_{i=1}^{n-1}f(x_i)\left[\bar{F}(x_i)\right]^{m_i}\right).$$
(1.1)

# **Key words and phrases:** Generalized order statistics, Novel family of distribution, single moments and product moments.

AMS (MOS) Subject Classifications: 62E10, 60E05. ISSN 1814-0432, 2025, http://ijmcs.future-in-tech.net Moreover, we assume that  $k \ge 1, n \ge 2$ , and  $\tilde{m} = (m_1, m_2, \cdots, m_{n-1}) \in \mathbb{R}^{n-1}$ holds the relationship  $\gamma_j = k + n - r + \sum_{j=r}^{n-1} m_j, 1 \le r \le n-1$  and  $\bar{F}(x) = 1 - F(x)$ . The marginal and joint density of GOS have been studied in two different contexts:  $\gamma_i \ne \gamma_j$  and  $\gamma_i = \gamma_j$ . Throughout the paper, we consider  $\gamma_i \ne \gamma_j$ . The PDF of  $r^{th} GOS$  is

$$f_{r,n,\tilde{m},k}(x) = C_{r-1}f(x)\sum_{i=1}^{r} a_i(r) \left[\bar{F}(x)\right]^{\gamma_i - 1}.$$
 (1.2)

The joint PDF of  $(r-s)^{th} GOS$  is

$$f_{r,s,n,\tilde{m},k}(x,y) = C_{s-1} \sum_{i=r+1}^{s} a_i^{(r)}(s) \left[\frac{\bar{F}(y)}{\bar{F}(x)}\right]^{\gamma_i} \sum_{i=1}^{r} a_i(r) \left[\bar{F}(x)\right]^{\gamma_i} \frac{f(x)}{\bar{F}(x)} \frac{f(y)}{\bar{F}(y)},$$
(1.3)

where

 $C_{r-1} = \prod_{i=1}^{r} \gamma_i, i = 1, 2, \cdots, n, a_i(r) = \prod_{j(\neq i)=1}^{r} \frac{1}{(\gamma_j - \gamma_i)}, \gamma_i \neq \gamma_j, 1 \le i \le r \le n \text{ and } a_i^{(r)}(s) = \prod_{j(\neq i)=r+1}^{s} \frac{1}{(\gamma_j - \gamma_i)}, \gamma_i \neq \gamma_j, r+1 \le i \le s \le n.$ The recurrence relation based on *GOS* is a long-standing practice and an

important issue in mathematical statistics just like others. In recent years, plenty of articles have dealt with the moments of generalized order statistics with several distributions. We refer the readers to [2, 3, 4, 5, 8, 9].

### 2 Kumaraswamy-Exponential Distribution

The Kumaraswamy-Exponential distribution (KED) is generated from the T - X family proposed by [7]. Suppose  $X_1, X_2, \dots, X_n$  are the KED RV with three parameters  $\alpha, \beta$  and  $\gamma$ . The CDF and PDF of KED are reported by

$$F(x) = 1 - \left[1 - (\alpha x)^{\beta}\right]^{\gamma}, \ 0 < x < 1, \text{ and } \alpha, \beta, \gamma > 0$$
 (2.4)

$$f(x) = \alpha^{\beta} \beta \gamma x^{\beta-1} \left[ 1 - (\alpha x)^{\beta} \right]^{\gamma-1}, \ x > 0,$$
(2.5)

respectively.

We note that the characterizing differential equation for KED from (2.5) and 2.4 is given by

$$\bar{F}(x) = \frac{1}{\alpha^{\beta}\beta\gamma} \left(x^{1-\beta} + \alpha^{\beta}x\right) f(x).$$
(2.6)

The *KED* has many applications in hydrology and industrial engineering and reduces to the Kumaraswamy distribution (KD) at  $\alpha = 1$ .

It has been aimed to establish the moment properties from the KED using GOS. Some recurrence relations for single and product moments are proved, which are not known in the existing literature.

# 3 Single Moments

The  $p^{th}$  moments of GOS for a random sample from F(x) is

$$\mu_{r,n,\tilde{m},k}^{p} = C_{r-1} \int_{-\infty}^{\infty} x^{p} f(x) \sum_{i=1}^{r} a_{i}(r) [\bar{F}(x)]^{\gamma_{i}-1} dx.$$

**Theorem 3.1.** Let  $X \sim KED(\alpha, \beta, \gamma)$  and  $n \in N$ ,  $k \ge 1$ . Then the following relation holds for single moments:

$$\mu_{r,n,\tilde{m},k}^p - \mu_{r-1,n,\tilde{m},k}^p = \frac{p}{\alpha^\beta \beta \gamma \gamma_r} \left[ \mu_{r,n,\tilde{m},k}^{p-\beta} - \alpha^\beta \mu_{r,n,\tilde{m},k}^p \right].$$
(3.7)

**Proof:** From [8], we have

$$\mu_{r,n,\tilde{m},k}^p - \mu_{r-1,n,\tilde{m},k}^p = pC_{r-2} \int_0^\infty x^{p-1} \sum_{i=1}^r a_i(r) [\bar{F}(x)]^{\gamma_i} dx.$$
(3.8)

Rewrite (3.8) as follows:

$$= \frac{p}{\alpha^{\beta}\beta\gamma} C_{r-2} \int_0^1 x^{p-1} \sum_{i=1}^r a_i(r) [\bar{F}(x)]^{\gamma_{i-1}} \left( x^{1-\beta} - \alpha^{\beta} x \right) f(x) dx.$$
(3.9)

Solving (3.9) confirms (3.7).

**Corollary 3.2.** Replacing p by -p in (3.7), we obtain the inverse moments of GOS for KED:

$$\mu_{r,n,\tilde{m},k}^{-p} - \mu_{r-1,n,\tilde{m},k}^{-p} = \frac{-p}{\alpha^{\beta}\beta\gamma\gamma_{r}} \left[ \mu_{r,n,\tilde{m},k}^{-(p+\beta)} - \alpha^{\beta}\mu_{r,n,\tilde{m},k}^{-p} \right].$$

**Corollary 3.3.** At  $\alpha = 1$ , Theorem 3.1 reduces to single moments of KD :

$$\mu_{r,n,\tilde{m},k}^p - \mu_{r-1,n,\tilde{m},k}^p = \frac{p}{\beta\gamma\gamma_r} \left[ \mu_{r,n,\tilde{m},k}^{p-\beta} - \mu_{r,n,\tilde{m},k}^p \right].$$

as obtained in [9].

**Remark 3.4.** (i) For order statistics (m = 0 and k = 1), the single moments from KED is  $E[X_{r:n}^p] - E[X_{r-1:n}^p] = \frac{p}{\alpha^\beta\beta\gamma\gamma_r}E[X_{r:n}^{p-\beta}] - \alpha^\beta E[X_{r:n}^p]$ .

(ii) For the k<sup>th</sup> record values 
$$(k = -1)$$
, the single moments from KED is  

$$E(X_n^{(k)})^p - E(X_{n-1}^{(k)})^p = \frac{p}{\alpha^\beta \beta \gamma k} \left[ E(Y_n^{(k)})^{p-\beta} - \alpha^\beta E(Y_n^{(k)})^p \right].$$

# 4 Product Moments

The  $(p,q)^{th}$  moments of GOS for a random sample from F(x) is

$$\mu_{r,s:n,\tilde{m},k}^{p,q} = \int_{-\infty}^{\infty} \int_{x_1}^{\infty} x_1^p x_2^q C_{s-1} \left\{ \sum_{i=r+1}^s a_i^{(r)}(s) \left( \frac{\bar{F}(x_2)}{\bar{F}(x_1)} \right)^{\gamma_i} \right\}$$
$$\times \sum_{i=1}^s a_i(r) [\bar{F}(x_1)]^{\gamma_i} \frac{f(x_1)}{\bar{F}(x_1)} \frac{f(x_2)}{(\bar{F}(x_2))} dx_2 dx_1.$$

**Theorem 4.1.** As stated in Theorem 3.1, the following relation holds for product moment

$$\mu_{r,s:n,\tilde{m},k}^{p,q} - \mu_{r,s-1:n,\tilde{m},k}^{p,q} = \frac{1}{\alpha^{\beta}\beta\gamma} \frac{q}{\gamma_s} \left[ \mu_{r,s:n,\tilde{m},k}^{p,q-\beta} - \alpha^{\beta} \mu_{r,s:n,\tilde{m},k}^{p,q} \right].$$
(4.10)

**Proof:** From [8], we have

$$\mu_{r,s:n,\tilde{m},k}^{p,q} - \mu_{r,s-1:n,\tilde{m},k}^{p,q} = qC_{s-2} \int_0^1 \int_{x_1}^1 x_1^p x_2^{q-1} \sum_{i=r+1}^s a_i^{(r)}(s) \left(\frac{\bar{F}(x_2)}{\bar{F}(x_1)}\right)^{\gamma_i} \\ \times \sum_{i=1}^s a_i(r) [\bar{F}(x_1)]^{\gamma_i} \frac{f(x_1)}{\bar{F}(x_1)} dx_2 dx_1.$$
(4.11)

$$= qC_{s-2} \int_{0}^{1} \int_{x_{1}}^{1} x_{1}^{p} x_{2}^{q-1} \sum_{i=r+1}^{s} a_{i}^{(r)}(s) \left(\frac{\bar{F}(x_{2})}{\bar{F}(x_{1})}\right)^{\gamma_{i}} \\ \times \sum_{i=1}^{s} a_{i}(r) [\bar{F}(x_{1})]^{\gamma_{i}} \frac{f(x_{1})}{\bar{F}(x_{1})} \frac{\bar{F}(x_{2})}{\bar{F}(x_{2})} dx_{2} dx_{1}.$$

$$(4.12)$$

Now, using (2.6) in (4.12), we obtain

$$\mu_{r:s,n,\tilde{m},k}^{p,q} - \mu_{r:s-1,n,\tilde{m},k}^{p,q} = qC_{s-2} \int_{-\infty}^{\infty} \int_{x_1}^{\infty} x_1^p x_2^{q-1} \sum_{i=r+1}^{s} a_i^{(r)}(s) \left(\frac{\bar{F}(x_2)}{\bar{F}(x_1)}\right)^{\gamma_i} \\ \times \sum_{i=1}^{s} a_i(r) [\bar{F}(x_1)]^{\gamma_i} \frac{f(x_1)}{\bar{F}(x_1)} \frac{f(x_2)}{\bar{F}(x_2)} \left\{ x_2^{1-\beta} - \alpha^{\beta} x_2 \right\} dx_2 dx_1,$$
(4.13)

and so  $\mu_{r,s:n,\tilde{m},k}^{p,q} = \mu_{r,s-1:n,\tilde{m},k}^{p,q} + \frac{1}{\alpha^{\beta}\beta\gamma} \frac{q}{\gamma_s} \left[ \mu_{r,s:n,\tilde{m},k}^{p,q-\beta} - \alpha^{\beta} \mu_{r,s:n,\tilde{m},k}^{p,q} \right].$ 

- **Remark 4.2.** (i) The relation for product moment based on order statistics for KED is  $\varphi_{r,s:n}^{p,q} - \varphi_{r,s-1:n}^{p,q} = \frac{1}{\alpha^{\beta}\beta\gamma} \frac{q}{(n-s+1)} \left[ \varphi_{r,s:n}^{p,q-\beta} - \alpha^{\beta}\varphi_{r,s:n}^{p,q} \right].$ 
  - (ii) The relation for product moment based on K th record values for KED is  $\varphi_{k(r,s)}^{p,q} \varphi_{k(r,s-1)}^{p,q} = \frac{1}{\alpha^{\beta}\beta\gamma} \frac{q}{k} \left[ \varphi_{k(r,s)}^{p,q-\beta} \alpha^{\beta}\varphi_{k(r,s)}^{p,q} \right].$

## References

- U. Kamps, A Concept of Generalized Order Statistics, B.G. Teubner, Stuttgart, 1995.
- [2] E. Cramer, U. Kamps, Relations for expectations of function of generalized order statistics, Journal of Statistical Planning and Inference, 89, (2000), 79–89.
- [3] A.A. Ahmad, Single and product moments of generalized order statistics from linear exponential distribution, Communications and Statistics-Theory and Methods, 37, (2008), 1162–1172.
- [4] M.I. Khan, M.A.R. Khan, Recurrence relations of moments of generalized order statistics in a compound Rayleigh distribution, Thailand Statistician, 15, no. 1, (2016), 11–16.
- [5] M.I. Khan, Generalized order statistics from Power Lomax distribution and characterization, Applied Mathematics and E-Notes, 18, (2018), 148–155.
- [6] A. Alzaatreh, F. Famoye, C. Lee, A new method for generating families of continuous distributions, Metron 71, (2013), 63–79.
- [7] P. Kumaraswamy, Generalized probability density-function for doublebounded random-processes, Journal of Hydrology, 46, (1980), 79–88.
- [8] H. Athar, H.M. Islam, Recurrence relations between single and product moments of generalized order statistics from a general class of distribution, Metron, 42, (2004), 327–337.
- [9] D. Kumar, Generalized Order Statistics from Kumaraswamy Distribution and its Characterization, Tamsui Oxford Journal of Information and Mathematical Sciences, 27, no. 4, (2011), 463–476.